

Sec 2.5 - Independent Random Variables

For continuous case

$$f_{2|1}(x_2 | x_1) = \frac{f_{1,2}(x_1, x_2)}{f_1(x_1)}$$

$$\Rightarrow f_{1,2}(x_1, x_2) = f_{2|1}(x_2 | x_1) \cdot f_1(x_1)$$

Suppose $f_{2|1}(x_2 | x_1)$ does not depend on x_1 .

Then

$$f_{2|1}(x_2 | x_1) = f_2(x_2)$$

Proof!

$$\begin{aligned}
 f_{2|1}(x_2 | x_1) &= f_{2|1}(x_2 | x_1) \underbrace{\int_{-\infty}^{\infty} f_1(x_1) dx_1}_{=1} \\
 &\stackrel{\text{doesn't depend on } x_1}{=} \int_{-\infty}^{\infty} \underbrace{f_{2|1}(x_2 | x_1) \cdot f_1(x_1)}_{f_{1,2}(x_1, x_2)} dx_1 \\
 &= \int_{-\infty}^{\infty} f_{1,2}(x_1, x_2) dx_1 \\
 &= f_2(x_2)
 \end{aligned}$$

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$$\therefore f_{2|1}(x_2|x_1) = f_2(x_2)$$

$$\text{and } f(x_1, x_2) = f_1(x_1) \cdot f_2(x_2)$$

def Independence

Let X_1, X_2 have joint pdf
 $f_{1,2}(x_1, x_2)$ [joint pmf $p_{1,2}(x_1, x_2)$]

and the marginal pdfs [pmfs]
 $f_1(x_1)$ [$p_1(x_1)$] and $f_2(x_2)$ [$p_2(x_2)$]

then random variables X_1 and X_2
are said to be independent

if and only if

$$f_{1,2}(x_1, x_2) = f_1(x_1) f_2(x_2)$$

$$p_{1,2}(x_1, x_2) = p_1(x_1) p_2(x_2).$$

If not independent, then they
are dependent.

ex) Assume X_1, X_2 are independent

$$\text{with } f_1(x_1) = \begin{cases} \frac{1}{2} x_1, & 0 < x_1 < 1 \\ 0, & \text{o.w.} \end{cases}$$

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$$f_2(x_2) = \begin{cases} \frac{1}{3} x_2^2, & 0 < x_2 < 1 \\ 0, & \text{o.w.} \end{cases}$$

then

$$f_{1,2}(x_1, x_2) = \begin{cases} \frac{1}{6} x_1 x_2^2, & 0 < x_1 < 1 \\ & 0 < x_2 < 1 \\ 0, & \text{o.w.} \end{cases}$$

Thm) X_1 and X_2 are independent

iff $f(x_1, x_2)$ can be written as a product of a nonnegative function of x_1 and a nonnegative function of x_2

$$f(x_1, x_2) \equiv g(x_1)h(x_2)$$

where $g(x_1) > 0$, $x_1 \in S_1$
and $h(x_2) > 0$, $x_2 \in S_2$

ex) $f(x_1, x_2) = \begin{cases} x_1 + x_2, & 0 < x_1 < 1 \\ & 0 < x_2 < 1 \\ 0, & \text{o.w.} \end{cases}$

Are x_1, x_2 independent?

$f(x_1) = \begin{cases} \int_0^1 (x_1 + x_2) dx_2 \\ = x_1 + \frac{1}{2}, & 0 < x_1 < 1 \\ 0, & \text{o.w.} \end{cases}$

$f(x_2) = \begin{cases} \int_0^1 (x_1 + x_2) dx_1 \\ = x_2 + \frac{1}{2}, & 0 < x_2 < 1 \\ 0, & \text{o.w.} \end{cases}$

Is $f(x_1, x_2) = f_1(x_1) \cdot f_2(x_2)$?
 $x_1 + x_2 \neq (x_1 + \frac{1}{2}) \cdot (x_2 + \frac{1}{2})$

NOT independent.

They are dependent.

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Thm! Let (X_1, X_2) have joint cdf $F(x_1, x_2)$ and let X_1 and X_2 have marginal cdfs $F(x_1)$ and $F(x_2)$
~~iff~~ X_1 and X_2 are independent
iff $F(x_1, x_2) = F_1(x_1) F_2(x_2)$
 $\forall (x_1, x_2) \in \mathbb{R}^2$.

Proof

(\Leftarrow) Suppose $F(x_1, x_2) = F(x_1)F(x_2)$.

then

$$f(x_1, x_2) = \frac{\partial^2}{\partial x_1 \partial x_2} F(x_1, x_2)$$

$$\text{i.e.} \quad \frac{\partial^2}{\partial x_1 \partial x_2} F(x_1)F(x_2)$$

$$= \left[\frac{\partial}{\partial x_1} F(x_1) \right] \left[\frac{\partial}{\partial x_2} F(x_2) \right]$$

$$= f(x_1) \cdot f(x_2)$$

$\therefore X_1$ and X_2 are independent.

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(\Rightarrow) Assume X_1, X_2 independent

$$F(x_1, x_2) = \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} f(w_1, w_2) dw_2 dw_1$$

$$\stackrel{\text{l.A.}}{=} \int_{-\infty}^{x_1} \int_{-\infty}^{x_2} f(w_1) f(w_2) dw_2 dw_1$$

$$= \left[\int_{-\infty}^{x_1} f(w_1) dw_1 \right] \left[\int_{-\infty}^{x_2} f(w_2) dw_2 \right]$$

$$= F(x_1) F(x_2) \quad \square$$

Thm! X_1 and X_2 are independent
iff

$$\mathbb{P}[a < X_1 \leq b, c < X_2 \leq d]$$

$$= \mathbb{P}[a < X_1 \leq b] \cdot \mathbb{P}[c < X_2 \leq d]$$

for every $\underbrace{a < b, c < d}_{\text{all constants}}$.

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thm ^{LP} X_1 and X_2 are independent
then $E[u(X_1)v(X_2)] = E[u(X_1)] \cdot E[v(X_2)]$

Proof.

Assume X_1, X_2 are independent.

$$E[u(X_1)v(X_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u(x_1)v(x_2) \cdot f_{1,2}(x_1, x_2) dx_1 dx_2$$

indep.

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u(x_1)v(x_2) f_1(x_1) f_2(x_2) dx_1 dx_2$$

$$= \left[\int_{-\infty}^{\infty} u(x_1) f_1(x_1) dx_1 \right] \left[\int_{-\infty}^{\infty} v(x_2) f_2(x_2) dx_2 \right]$$

$$= E[u(X_1)] \cdot E[v(X_2)]$$

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Note) If X, Y are independent
then $\text{Cov}(X, Y) = 0$

Proof)

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}[(X - \mu_x)(Y - \mu_y)] \\ &\stackrel{\text{indep}}{=} \mathbb{E}[X - \mu_x] \mathbb{E}[Y - \mu_y] \\ &= \left\{ \mathbb{E}[X] - \underbrace{\mathbb{E}[X]}_{\mu_x} \right\} \left\{ \mathbb{E}[Y] - \underbrace{\mathbb{E}[Y]}_{\mu_y} \right\} \\ &= 0\end{aligned}$$

Converse not generally true.

ex) Let (X, Y) have joint pmf.

$$P_{X,Y}(x,y) = \begin{cases} \frac{1}{8} & , \quad \begin{array}{l} x = -1, 0, 1 \\ y = -1, 0, 1 \\ \text{but } (x,y) \neq (0,0) \end{array} \\ 0, & \text{o.w.} \end{cases}$$

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		y			
		-1	0	1	p(x)
x	-1	1/8	1/8	1/8	3/8
	0	1/8	0	1/8	2/8
	1	1/8	1/8	1/8	3/8
	p(y)	3/8	2/8	3/8	1

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 0$$

$$\begin{aligned}
 E[XY] &= \sum_x \sum_y xy p(x, y) = 0 \\
 &= -1(-1)\left(\frac{1}{8}\right) + (-1)(1)\left(\frac{1}{8}\right) \\
 &\quad + 0 + 0 + 0 \\
 &\quad + 1(-1)\left(\frac{1}{8}\right) + 0 + 1(1)\left(\frac{1}{8}\right)
 \end{aligned}$$

$$\begin{aligned}
 E[X] &= -1\left(\frac{3}{8}\right) + 0\left(\frac{2}{8}\right) + 1\left(\frac{3}{8}\right) = 0 \\
 &= E[Y]
 \end{aligned}$$

Indep? $P(x, y) = P(x)P(y)$

$$\begin{aligned}
 P\left[X=1, Y=1\right] &= \frac{1}{8} \stackrel{?}{=} P[X=1]P[Y=1] \\
 &= \frac{1}{8} \neq \frac{3}{8} \cdot \frac{3}{8}
 \end{aligned}$$

Dependent

Thm Suppose the joint MGF $M(t_1, t_2)$ exists for RVs X_1, X_2 .

then X_1 and X_2 are independent iff

$$M(t_1, t_2) = M(t_1, 0) M(0, t_2)$$

i.e. you can write the joint MGF as a product of the marginal MGFs.

ex! in sec 2.2

$$P_{j2}(x_1, x_2) = \begin{cases} \frac{\mu_1^{x_1} \mu_2^{x_2} e^{-\mu_1} e^{-\mu_2}}{x_1! x_2!} & / \begin{matrix} x_1=0,1,2,\dots \\ x_2=0,1,2,\dots \end{matrix} \\ 0 & / \text{o.w.} \end{cases}$$

$$M(t_1, t_2) = E[e^{t_1 X_1 + t_2 X_2}] = \sum_{x_1=0}^{\infty} \sum_{x_2=0}^{\infty} e^{t_1 x_1 + t_2 x_2} \frac{\mu_1^{x_1} \mu_2^{x_2} e^{-\mu_1} e^{-\mu_2}}{x_1! x_2!}$$

$$= e^{\mu_1(e^{t_1}-1)} e^{\mu_2(e^{t_2}-1)}$$

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$$M(t_1, 0) = \mathbb{E}[e^{t_1 X_1}]$$

$$= \sum_{x_1=0}^{\infty} \sum_{x_2=0}^{\infty} \frac{e^{t_1 x_1} \mu_1^{x_1} \mu_2^{x_2} e^{-\mu_1} e^{-\mu_2}}{x_1! x_2!}$$

$$= \sum_{x_1=0}^{\infty} \frac{e^{t_1 x_1} \mu_1^{x_1} e^{-\mu_1}}{x_1!} \underbrace{\sum_{x_2=0}^{\infty} \frac{\mu_2^{x_2} e^{-\mu_2}}{x_2!}}_{=1}$$

$$= e^{-\mu_1} \sum_{x_1=0}^{\infty} \frac{e^{t_1 x_1} \mu_1^{x_1}}{x_1!}$$

$$= e^{-\mu_1} \sum_{x_1=0}^{\infty} \frac{(e^{t_1} \mu_1)^{x_1}}{x_1!}$$

$$= e^{-\mu_1} e^{\mu_1 e^{t_1}} = e^{\mu_1 (e^{t_1} - 1)}$$

since $\sum_{k=0}^{\infty} \frac{a^k}{k!} = e^a$

$$\text{Similarly } M(0, t_2) = e^{\mu_2 (e^{t_2} - 1)}$$

$\therefore M(t_1, t_2) = M(t_1, 0) M(0, t_2)$ and X_1, X_2 are independent.