

Stat 401 - 12/8/17

①

(21)  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} \text{Unif}(0, \theta)$ .  
↑  
unknown.

Let  $Y_n = \max \{X_1, \dots, X_n\}$ .

$$a) F_{Y_n}(t) = \begin{cases} 1, & t > \theta \\ \left(\frac{t}{\theta}\right)^n, & 0 < t \leq \theta \\ 0, & t \leq 0. \end{cases}$$

If  $X_i \sim \text{Unif}(0, \theta)$ .

$$f(x) = \frac{1}{\theta - 0} = \frac{1}{\theta}, \quad 0 < x < \theta$$

$$\mathbb{P}[Y_n \leq t] = \mathbb{P}[X_1 \leq t, X_2 \leq t, \dots, X_n \leq t]$$

WLOG Let  $Y_n = X_i$

↑  
largest of  $X_i$ 's.

indep

$$\rightarrow = \mathbb{P}[X_1 \leq t] \cdots \mathbb{P}[X_n \leq t]$$

$$= \left(\mathbb{P}[X_1 \leq t]\right)^n$$

What is  $P[X \leq t]$

$$= \int_0^t \frac{1}{\theta} dt$$

$$= \left. \frac{t}{\theta} \right|_0^t = \frac{t}{\theta} - \frac{0}{\theta} = \frac{t}{\theta}$$

→  $\left( P[X_1 \leq t] \right)^n = \left( \frac{t}{\theta} \right)^n$   
for  $0 < t < \theta$

b) Find the PDF for  $Y_n$ .

$$f_{Y_n}(t) = \begin{cases} \frac{n}{\theta^n} \cdot t^{n-1}, & 0 < t < \theta \\ 0, & \text{o.w.} \end{cases}$$

c) Show  $Y_n$  is a biased estimator of  $\theta$ .

Recall unbiased est means  $E[Y_n] = \theta$

If  $E[Y_n] \neq \theta$  then biased.

$$E(X) = \int x f(x) dx$$

(3)

$$E[Y_n] = \int_0^{\theta} t \cdot \frac{n}{\theta^n} t^{n-1} dt$$

$$= \frac{n}{\theta^n} \int_0^{\theta} t^n dt$$

$$= \frac{n}{\theta^n} \left. \frac{t^{n+1}}{n+1} \right|_0^{\theta}$$

$$= \frac{n}{\theta^n} \frac{\theta^{n+1}}{n+1} - \frac{n}{\theta^n} \cdot 0$$

$$= \frac{n}{n+1} \theta \neq \theta$$

$\therefore$  Biased.

d) Show  $\frac{n+1}{n} Y_n$  is unbiased for  $\theta$ .

Show  $E\left[\frac{n+1}{n} Y_n\right] = \theta$ .

$$E\left[\frac{n+1}{n} Y_n\right] = \frac{n+1}{n} E[Y_n]$$

$$= \frac{n+1}{n} \cdot \frac{n}{n+1} \theta = \theta \quad \checkmark$$



e) Show  $Y_n \xrightarrow{P} \theta$

Show  $\lim_{n \rightarrow \infty} P \{ |Y_n - \theta| > \varepsilon \} = 0$

$$P \{ |Y_n - \theta| > \varepsilon \}$$

$$= P \{ \theta - Y_n > \varepsilon \}$$

we have  $0 < t < \theta$

ex  $0 < t < 2$   
 $Y_n = 1$

$$|Y_n - \theta| = |1 - 2| = |-1| = 1$$

ex If use  $A$  instead of  $Y$ .

$$0 < a < \theta$$

$$A = \begin{cases} \frac{n}{\theta^n} a^{n-1}, & 0 < a < \theta \\ 0, & \text{o.w.} \end{cases}$$



(5)

$$= \mathbb{P} \left\{ -Y_n > \varepsilon - \theta \right\}$$

$$= \mathbb{P} \left\{ Y_n < \theta - \varepsilon \right\}$$

$$= F_{Y_n}(\theta - \varepsilon)$$

$$= \left( \frac{\theta - \varepsilon}{\theta} \right)^n$$

$$= \left( \cancel{\theta} \left( 1 - \frac{\varepsilon}{\theta} \right) \right)^n$$

Since  $\varepsilon > 0$  it is "small"

WLOG assume  $0 < \varepsilon < \theta$

$$\Rightarrow 0 < \frac{\varepsilon}{\theta} < 1$$

$$\Rightarrow 1 - \frac{\varepsilon}{\theta} < 1$$

$$\rightarrow \lim_{n \rightarrow \infty} \left( 1 - \frac{\varepsilon}{\theta} \right)^n = 0$$

$$\therefore \mathbb{P} \left\{ |Y_n - \theta| > \varepsilon \right\} \rightarrow 0$$

as  $n \rightarrow \infty$ .

$$\therefore Y_n \xrightarrow{P} \theta$$

(6)

f) Show  $\frac{n+1}{n} Y_n \xrightarrow{P} \theta$ .

$$0 \leq \mathbb{P} \left\{ \left| \frac{n+1}{n} Y_n - \theta \right| > \varepsilon \right\}$$

$$= \mathbb{P} \left\{ \frac{n+1}{n} \left| Y_n - \frac{n}{n+1} \theta \right| > \varepsilon \right\}$$

$$= \mathbb{P} \left\{ \left| Y_n - \frac{n}{n+1} \theta \right| > \frac{n}{n+1} \varepsilon \right\}.$$

$\mathbb{E}[Y_n]$

Chebyshev's Ineq.

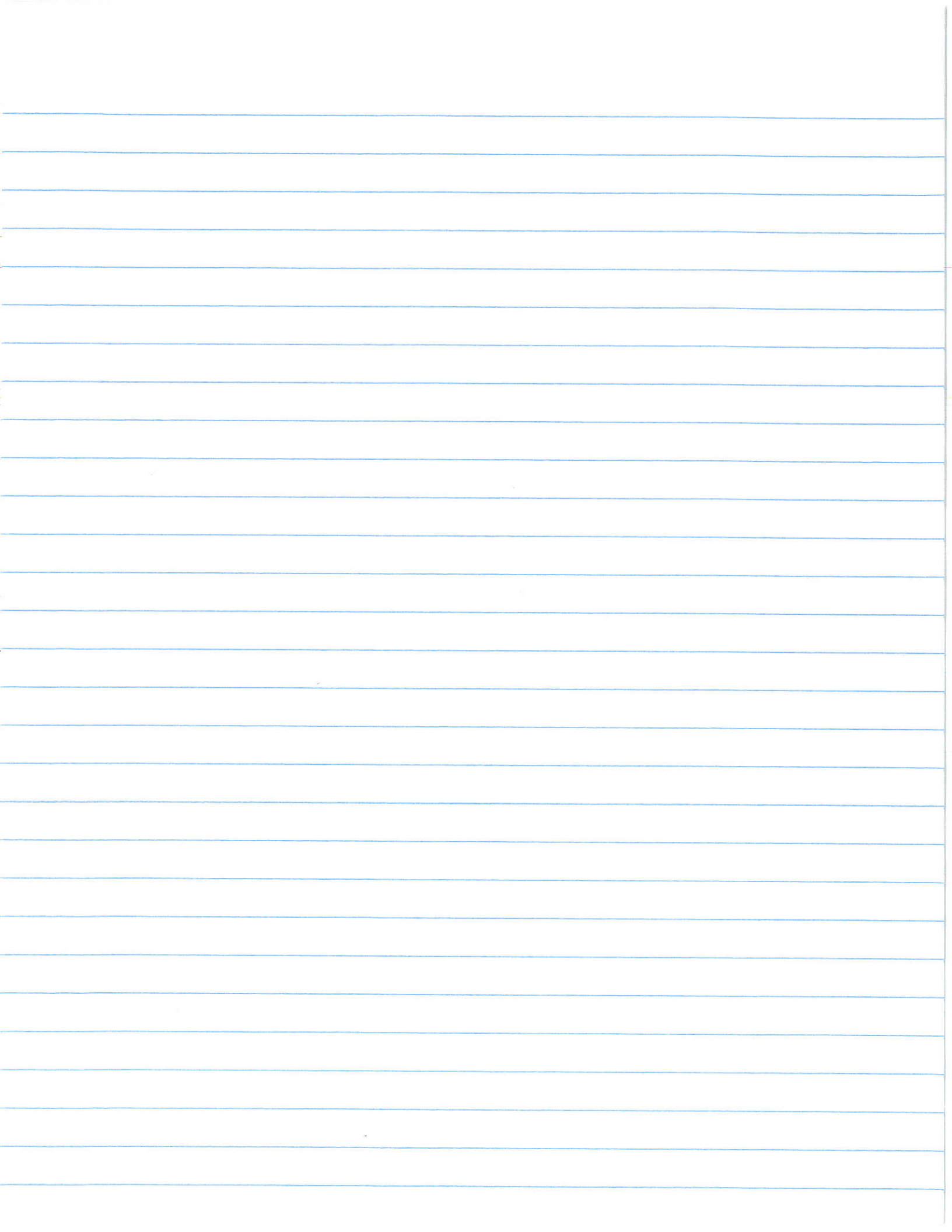
$$\mathbb{P} \left\{ |X - \mu| > \varepsilon \right\} \leq \frac{\text{Var}[X]}{\varepsilon^2}$$

$\downarrow$   
 $\mathbb{E}[X]$

$$\leq \frac{\text{Var}[Y_n]}{\left(\frac{n}{n+1} \varepsilon\right)^2} = \frac{\left[\frac{n}{n+2} - \frac{n^2}{(n+1)^2}\right] \theta^2}{\frac{n^2}{(n+1)^2} \varepsilon^2}$$

$\dots \rightarrow 0$  as  $n \rightarrow \infty$ .





(23) Consider  $Z_n = n(\theta - Y_n)$

Let  $t \in (0, n\theta)$

Show  $Z_n \xrightarrow{D} \text{Exp}(\theta) = Z$

One method:

Show  $\lim_{n \rightarrow \infty} F_{Z_n}(t) = F_Z(t)$

$$\begin{aligned}
 P[Z_n \leq t] &= P[n(\theta - Y_n) \leq t] \\
 &= P\left[\theta - Y_n \leq \frac{t}{n}\right] \\
 &= P\left[-Y_n \leq \frac{t}{n} - \theta\right] \\
 &= P\left[Y_n \geq \theta - \frac{t}{n}\right] \\
 &= 1 - P\left(Y_n \leq \theta - \frac{t}{n}\right) \\
 &= 1 - F_{Y_n}\left(\theta - \frac{t}{n}\right) \\
 &= 1 - \left(\frac{\theta - \frac{t}{n}}{\theta}\right)^n \\
 &= 1 - \left(1 - \frac{t}{n\theta}\right)^n
 \end{aligned}$$



Monday 12-3; Thurs 9-2 or 3 (8)

$$= 1 - \left(1 + \frac{-t/\theta}{n}\right)^n$$

Recall

$$\lim_{n \rightarrow \infty} \left[1 + \frac{b}{n} + \frac{\psi(n)}{n}\right]^{cn}$$

$$= \lim_{n \rightarrow \infty} \left[1 + \frac{b}{n}\right]^{cn}$$

$$= e^{bc}$$

$$b = -t/\theta \quad c = 1.$$

$$\lim_{n \rightarrow \infty} \mathbb{P}[Z_n \leq t]$$

$$= \lim_{n \rightarrow \infty} 1 - \left(1 + \frac{-t/\theta}{n}\right)^n$$

$$= 1 - e^{-\frac{t}{\theta}}$$

← CDF for  
exponential  
RV  
w/ parameter  
 $\theta$ .

$$\therefore Z_n \xrightarrow{D} Z = \text{Exp}(\theta)$$