

Name: _____ UIN: _____

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2. In order to receive full credit for a problem, you should show all of your work and explain your reasoning. Good work can receive substantial partial credit even if the final answer is incorrect.
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Problem 1. [30 points]

Suppose X follow an uniform distribution with parameter $\theta > 0$, i.e. $X \sim \text{Unif}(0, \theta)$, its density function is

$$f(x; \theta) = \frac{1}{\theta}, 0 < x \leq \theta.$$

A random sample $X_1, \dots, X_n$ is drawn from the distribution.

- (1). [6 points]. Find its cumulative distribution function $F(x)$.

$$\forall 0 < x \leq \theta, \quad F(x) = \int_0^x \frac{1}{\theta} dx = \frac{x}{\theta}$$

$$\therefore \text{cdf } F(x) = \begin{cases} 0 & x \leq 0 \\ \frac{x}{\theta} & 0 < x \leq \theta \\ 1 & x > \theta \end{cases}$$

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Total		

- (2). [10 points]. Let $X_{(n)}$ be the largest order statistics of the sample, find its density distribution.

Distribution Function of $Y = X_{(n)}$, $Y = (0, \theta]$

$$\begin{aligned} \forall 0 < y \leq \theta, \quad F_Y(y) &= P(Y \leq y) = P(X_{(n)} \leq y) \\ &= P(X_1, \dots, X_n \leq y) = \prod_{i=1}^n P(X_i \leq y) \\ &= [F(y)]^n = \left(\frac{y}{\theta}\right)^n \end{aligned}$$

Density function

$$f(y) = F'(y) = n \left(\frac{y}{\theta}\right)^{n-1} \cdot \left(\frac{1}{\theta}\right) = \frac{n y^{n-1}}{\theta^n}, 0 < y \leq \theta.$$

- (3). [8 points] Find constant c such that $cX_{(n)}$ is unbiased for parameter θ .

$$\begin{aligned}\theta &= E[cX_{(n)}] = c \cdot EX_{(n)} = c \int_0^\theta y \cdot \frac{n y^{n-1}}{\theta^n} dy \\ &= \frac{n \cdot c}{\theta^n} \int_0^\theta y^n dy = \frac{n \cdot c}{\theta^n} \cdot \frac{y^{n+1}}{n+1} \Big|_0^\theta \\ &= \frac{n}{n+1} \theta \cdot c \\ \Rightarrow c &= \frac{n+1}{n}\end{aligned}$$

- (4). [6 points] Find the limiting distribution of $Z_n = n(\theta - X_{(n)})$.

$$S_Y = (0, \theta], \quad S_{Z_n} = [0, n\theta), \quad Z_n = n(\theta - Y)$$

$$\begin{aligned}\text{CDF } F_{Z_n}(z) &= P[n(\theta - Y) \leq z] \\ &= P[Y \geq \theta - \frac{z}{n}] \\ &= 1 - P(Y < \theta - \frac{z}{n}) \\ &= 1 - F(\theta - \frac{z}{n}) = 1 - \left(\frac{\theta - \frac{z}{n}}{\theta}\right)^n \\ &= 1 - \left(1 - \frac{z/\theta}{n}\right)^n \\ &\xrightarrow{\text{as } n \rightarrow \infty} 1 - e^{-z/\theta} = F_Z(z),\end{aligned}$$

where $Z \sim \text{exponential distribution } (\theta)$.
 $S_{Z_n} \rightarrow S_Z = (0, +\infty)$

Problem 2. [30 points]

Suppose X follows an exponential distribution, i.e. $X \sim \text{Exp}(1/\theta)$

$$f(x) = \theta e^{-\theta x}, x > 0. \quad \sim T(1, \frac{1}{\theta})$$

- (1). [8 points] Calculate its moment generating function $M_X(t)$.

$$\begin{aligned}M_X(t) &= E(e^{tX}) = \int_0^\infty e^{tx} \cdot \theta e^{-\theta x} dx \\ &= \theta \int_0^\infty e^{(t-\theta)x} dx \\ &= \frac{\theta}{-(t-\theta)} \int_0^\infty e^{-v} dv \quad \text{let } v = (t-\theta)x \\ &= \frac{\theta}{\theta-t} = \left(1 - \frac{t}{\theta}\right)^{-1}\end{aligned}$$

- (2). [8 points] Calculate its mean $E(X)$ and variance $\text{Var}(X)$.

[Gamma function: $\Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx$.]

$$\begin{aligned}E(X) &= \int_0^\infty x \theta e^{-\theta x} dx = \frac{1}{\theta} \int_0^\infty v e^{-v} dv = \frac{1}{\theta} = \frac{1}{\theta^{(1)}} \\ E(X^2) &= \int_0^\infty x^2 \theta e^{-\theta x} dx = \frac{1}{\theta^2} \int_0^\infty v^2 e^{-v} dv = \frac{2}{\theta^2} = \frac{2}{\theta^{(2)}} \\ \text{Var } X &= E(X^2) - (E(X))^2 = \frac{2}{\theta^2} - \frac{1}{\theta^2} = \frac{1}{\theta^2}\end{aligned}$$

$$\begin{aligned}(2) \quad E(X) &= M'(0) = \frac{\partial}{\partial t} \left(\frac{\theta}{\theta-t} \right) \Big|_{t=0} = \frac{1}{\theta^2} \\ E(X^2) &= M''(0) = \frac{\partial^2}{\partial t^2} \left(\frac{\theta}{\theta-t} \right) \Big|_{t=0} = \frac{2}{\theta^2} \\ \text{Var } X &= E(X^2) - (E(X))^2 = \frac{1}{\theta^2}\end{aligned}$$

(3). [6 points] A random sample $X_1, \dots, X_n$ is drawn from the distribution above, what is the distribution of $Y_n = X_1 + \dots + X_n$?

[The MGF of Gamma distribution $\Gamma(\alpha, \beta)$, $M(t) = (1 - \beta t)^{-\alpha}$.]

$$M_{Y_n}(t) = E(e^{tY_n}) = E(e^{t(X_1 + \dots + X_n)})$$

independent

$$= \prod_{i=1}^n E(e^{tX_i})$$

$$= \left[1 - \frac{t}{\theta}\right]^{-n}$$

$$\therefore Y_n \sim \mathcal{P}(n, \frac{1}{\theta})$$

(4). [8 points] Show that Y_n/n converges in probability to $1/\theta$.

$$\frac{Y_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X}, \quad \text{Var}(\bar{X}) = \frac{\text{Var}(X)}{n} = \frac{1}{n\theta^2}$$

X : mean $\frac{1}{\theta}$, variance $\frac{1}{\theta^2}$

$$\mathbb{P}(|\bar{X} - \frac{1}{\theta}| > \varepsilon) \leq \frac{\text{Var}(\bar{X})}{\varepsilon^2} = \frac{1}{n\theta^2\varepsilon^2} \rightarrow 0$$

i.e. $\bar{X} \xrightarrow{P} \frac{1}{\theta}$

$$\textcircled{2} \text{ By WLLN: } \bar{X} = \frac{Y_n}{n} \xrightarrow{P} EX = \frac{1}{\theta}$$

3. [40 points]. Consider a random sample $(X_1, \dots, X_n)$ drawn from a normal population with mean θ and variance σ^2 (known)

$$f(x) = (2\pi\sigma^2)^{-1/2} \exp\left\{-\frac{(x-\theta)^2}{2\sigma^2}\right\}.$$

(1). [12 points] Find the mle $\hat{\theta}$ for θ .

$$L(\theta) = \prod_{i=1}^n f(x_i) = (2\pi\sigma^2)^{-n/2} \cdot \exp\left\{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2\right\}$$

$$L(\theta) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2$$

$$\frac{\partial L(\theta)}{\partial \theta} = 0 \Rightarrow \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta) = 0, \quad \frac{\partial^2 L(\theta)}{\partial \theta^2} = -\frac{n}{\sigma^2} < 0$$

$$\Rightarrow \hat{\theta}_{MLE} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{X}$$

$E\bar{X} = \theta$ unbiased

(2). [12 points] Check if the mle $\hat{\theta}$ is efficient for θ .

$$\frac{\partial \log f(x, \theta)}{\partial \theta} = \frac{\partial}{\partial \theta} \left[-\frac{1}{2} \log(2\pi\sigma^2) - \frac{(x-\theta)^2}{2\sigma^2} \right] = \frac{2(x-\theta)}{2\sigma^2} = \frac{x-\theta}{\sigma^2}$$

$$\frac{\partial^2 \log f(x, \theta)}{\partial \theta^2} = \frac{\partial}{\partial \theta} \left[\frac{x-\theta}{\sigma^2} \right] = -\frac{1}{\sigma^2}$$

$$\text{Fisher Information } I(\theta) = -E\left[\frac{\partial^2 \log f(x, \theta)}{\partial \theta^2}\right] = \frac{1}{\sigma^2}$$

$$\text{Var}(\hat{\theta}) = \text{Var}(\bar{X}) = \frac{\sigma^2}{n} \quad \left. \vphantom{\text{Var}(\hat{\theta})} \right\} \quad \text{Var}(\hat{\theta}) = \frac{[I(\theta)]^{-1}}{nI(\theta)} = \frac{1}{nI(\theta)}$$

$$\frac{1}{nI(\theta)} = \frac{1}{n \cdot \frac{1}{\sigma^2}} = \frac{\sigma^2}{n}$$

$$K(\theta) = E(\bar{X}) = \theta \quad \text{unbiased} \quad \text{Var}(\hat{\theta}) = \frac{1}{nI(\theta)}$$

$\therefore \hat{\theta}$ is an efficient estimator for θ .

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- (4). [10 points] What is the sampling distribution of the estimator? Construct a 100(1 - α)% confidence interval for the unknown mean μ .

$$X_1, \dots, X_n \sim N(\theta, \sigma^2)$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim N\left(\frac{1}{n} \sum_{i=1}^n E(X_i), \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i)\right)$$

$$\text{i.e., } \bar{X} \sim N\left(\theta, \frac{\sigma^2}{n}\right)$$

$$P\left(-Z_{\alpha/2} < \frac{\bar{X} - \theta}{\sqrt{\sigma^2/n}} < Z_{\alpha/2}\right) = 1 - \alpha$$

$$\text{C.I. for } \theta: \bar{X} \pm Z_{\alpha/2} \cdot \frac{\sigma}{\sqrt{n}}$$

- (5). [6 points] For a new parameter $\eta = \mu^3$, find its mle and its asymptotic distribution.

Based on functional invariant property of mle,

$$\hat{\eta}_{\text{mle}} = (\hat{\theta}_{\text{mle}})^3 = \bar{X}^3$$

Delta method $g(\theta) = (\theta^3)' = 3\theta^2$

$$\sqrt{n}(\hat{\eta}_{\text{mle}} - \eta) \rightsquigarrow N\left(0, \frac{[g'(\theta)]^2}{I(\theta)}\right)$$

$$\rightsquigarrow N(0, 9\theta^4)$$