

Analysis of Variance - RCBD

(1) Parameter Estimator: least square / Maximum Likelihood

$$\text{Model: } Y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}, \quad \varepsilon_{ij} \stackrel{\text{iid}}{\sim} N(0, \sigma^2) \\ i=1, \dots, k, \quad j=1, \dots, b \\ \sum_{i=1}^k \tau_i = 0, \quad \sum_{j=1}^b \beta_j = 0 \quad (\text{both fixed factors})$$

$$\Rightarrow \hat{\mu} = \bar{Y}_{..}, \quad \hat{\tau}_i = \bar{Y}_{i.} - \bar{Y}_{..}, \quad \hat{\beta}_j = \bar{Y}_{.j} - \bar{Y}_{..}$$

$$\text{Then fitted value: } \hat{Y}_{ij} = \hat{\mu}_{ij} = \hat{\mu} + \hat{\tau}_i + \hat{\beta}_j$$

$$\hat{Y}_{ij} = \bar{Y}_{i.} + \bar{Y}_{.j} - \bar{Y}_{..}$$

$$\text{Residual: } e_{ij} = Y_{ij} - \hat{Y}_{ij} = Y_{ij} - \bar{Y}_{i.} - \bar{Y}_{.j} + \bar{Y}_{..}$$

(2) S.S. Decomposition

$$\text{Note that: } \sum_i (\bar{Y}_{i.} - \bar{Y}_{..}) = 0, \quad \sum_j (\bar{Y}_{.j} - \bar{Y}_{..}) = 0$$

$$\sum_j (\bar{Y}_{ij} - \bar{Y}_{i.}) = 0 \quad \forall i=1, \dots, k \quad \Rightarrow \sum_j e_{ij} = 0, \quad \forall i=1, \dots, k$$

$$\sum_i (\bar{Y}_{ij} - \bar{Y}_{.j}) = 0 \quad \forall j=1, \dots, b \quad \Rightarrow \sum_i e_{ij} = 0, \quad \forall j=1, \dots, b$$

$$\sum_i \sum_j e_{ij} = \sum_i \sum_j (Y_{ij} - \bar{Y}_{i.} - \bar{Y}_{.j} + \bar{Y}_{..}) = 0$$

$$\left[\sum_i \sum_j Y_{ij} = b \sum_i \bar{Y}_{i.} = k \cdot \sum_j \bar{Y}_{.j} = bk \cdot \bar{Y}_{..} \right]$$

Deviation Decomposition

$$Y_{ij} - \bar{Y}_{..} = (\bar{Y}_{i.} - \bar{Y}_{..}) + (\bar{Y}_{.j} - \bar{Y}_{..}) + (Y_{ij} - \bar{Y}_{i.} - \bar{Y}_{.j} + \bar{Y}_{..})$$

$$Y_{ij} - \bar{Y}_{..} = \underset{(\hat{\tau}_i)}{A_i} + \underset{(\hat{\beta}_j)}{B_j} + e_{ij}$$

$$\begin{aligned}\sum_i \sum_j (Y_{ij} - \bar{Y}_{..})^2 &= \sum_i \sum_j (A_i + B_j + e_{ij})^2 \\ &= \sum_i \sum_j A_i^2 + \sum_i \sum_j B_j^2 + \sum_i \sum_j e_{ij}^2 \\ &\quad + 2 \left(\sum_i \sum_j A_i e_{ij} + \sum_i \sum_j B_j e_{ij} + \sum_i \sum_j A_i B_j \right)\end{aligned}$$

It can be shown that

$$\sum_i \sum_j A_i e_{ij} = \sum_i A_i \cdot \left(\sum_j e_{ij} \right) = 0$$

$$\sum_i \sum_j B_j e_{ij} = \sum_j B_j \cdot \left(\sum_i e_{ij} \right) = 0$$

$$\sum_i \sum_j A_i B_j = \left(\sum_i A_i \right) \left(\sum_j B_j \right) = 0$$

$$\sum_{i=1}^k \sum_{j=1}^b (Y_{ij} - \bar{Y}_{..})^2 = \sum_{i=1}^k \sum_{j=1}^b (\bar{Y}_{i.} - \bar{Y}_{..})^2 + \sum_{i=1}^k \sum_{j=1}^b (\bar{Y}_{.j} - \bar{Y}_{..})^2 + \sum_{i=1}^k \sum_{j=1}^b e_{ij}^2$$

$$\sum_{i=1}^k \sum_{j=1}^b (Y_{ij} - \bar{Y}_{..})^2 = b \sum_{i=1}^k (\bar{Y}_{i.} - \bar{Y}_{..})^2 + k \sum_{j=1}^b (\bar{Y}_{.j} - \bar{Y}_{..})^2 + \sum_{i=1}^k \sum_{j=1}^b e_{ij}^2$$

$$SSTO = \overset{\downarrow}{SSR} + \overset{\downarrow}{SSB} + \overset{\nwarrow}{SSE}$$

And $DF(SSTO) = kb - 1 = 0$ restriction $\sum_i \sum_j (Y_{ij} - \bar{Y}_{..}) = 0$

$DF(SSB) = b - 1$ $\sum_j (\bar{Y}_{.j} - \bar{Y}_{..}) = 0$

$DF(SSR) = k - 1$ $\sum_i (\bar{Y}_{i.} - \bar{Y}_{..}) = 0$

$DF(SSE) = (b-1)(k-1)$ $\begin{cases} \sum_i e_{ij} = 0 & \forall j=1, \dots, b \\ \sum_j e_{ij} = 0 & \forall i=1, \dots, k \end{cases}$

Based on Cochran's Theorem,

$\frac{SSTR}{\sigma^2}$, $\frac{SSB}{\sigma^2}$, $\frac{SSE}{\sigma^2}$ are independent χ^2 random variables.

i.e. $\frac{SSTR}{\sigma^2} \sim \chi^2(k-1)$, $\frac{SSB}{\sigma^2} \sim \chi^2(b-1)$, $\frac{SSE}{\sigma^2} \sim \chi^2((k-1)(b-1))$

and $SSTR \perp\!\!\!\perp SSB \perp\!\!\!\perp SSE$ (mutually independent)

Under $H_0: \tau_1 = \dots = \tau_k = 0$, $H_1: \text{not all } \tau_i = 0$

$$F_{TR} = \frac{\cancel{SSTR}}{MSE} \sim F(k-1, (k-1)(b-1))$$

Under $H_0: \beta_1 = \dots = \beta_b = 0$, $H_1: \text{not all } \beta_j = 0$

$$F_B = \frac{MSB}{MSE} \sim F(b-1, (k-1)(b-1))$$

And $E[MSE] = E\left[\frac{SSE}{(k-1)(b-1)}\right] = \sigma^2$

$$E[MSTR] = E\left[\frac{SSTR}{k-1}\right] = \sigma^2 + b \cdot \left(\frac{1}{k-1} \sum_{i=1}^k \tau_i^2\right)$$

$$E[MSB] = E\left[\frac{SSB}{b-1}\right] = \sigma^2 + k \cdot \left(\frac{1}{b-1} \sum_{j=1}^b \beta_j^2\right)$$

$$E[MSTR] = E\left[\frac{SSTR}{k-1}\right] = E\left[\frac{1}{k-1} \left(b \cdot \sum_{i=1}^k (\bar{Y}_{i.} - \bar{Y}_{..})^2\right)\right]$$

$$Y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}, \quad \varepsilon_{ij} \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

$$\Rightarrow Y_{ij} \stackrel{\text{ind.}}{\sim} N(\mu + \tau_i + \beta_j, \sigma^2)$$

Treatment mean at i -th level

$$\bar{Y}_{i.} = \frac{1}{b} \sum_{j=1}^b Y_{ij}$$

$$\bar{Y}_{i.} \stackrel{\text{ind.}}{\sim} N(\mu + \tau_i, \frac{\sigma^2}{b})$$

$$\begin{cases} \text{mean} & E(\bar{Y}_{i.}) = \frac{1}{b} \sum_{j=1}^b (\mu + \tau_i + \beta_j) \stackrel{\sum \beta_j = 0}{=} \mu + \tau_i \\ \text{variance} & \text{Var}(\bar{Y}_{i.}) \stackrel{\text{ind.}}{=} \frac{1}{b^2} \sum_{j=1}^b \text{Var}(Y_{ij}) = \frac{\sigma^2}{b} \end{cases}$$

$$\text{Under } H_0: \tau_1 = \dots = \tau_k = 0, \quad \bar{Y}_{i.} \stackrel{H_0}{\sim} N(\mu, \frac{\sigma^2}{b})$$

$$\text{Grand mean } \bar{Y}_{..} = \frac{1}{kb} \sum_i \sum_j Y_{ij} = \frac{1}{k} \sum_{i=1}^k \bar{Y}_{i.} \sim N(\mu, \frac{\sigma^2}{kb})$$

Under H_0 : $\bar{Y}_{..}$ is ~~an~~ average of $\bar{Y}_{i.}$ (k independent observations) with same mean.

Following Student's t -theorem

$$\frac{\sum_{i=1}^k (\bar{Y}_{i.} - \bar{Y}_{..})^2}{(\frac{\sigma^2}{b})} \sim \chi^2(k-1)$$

$$\begin{cases} X_i = \bar{Y}_{i.} & \bar{X} = \bar{Y}_{..} \\ \text{Var}(X_i) = \frac{\sigma^2}{b}, \quad i=1, \dots, k \end{cases}$$

$$E\left\{\frac{1}{k-1} \left[b \cdot \sum_{i=1}^k (\bar{Y}_{i.} - \bar{Y}_{..})^2\right]\right\} \stackrel{H_0: \tau_i=0 \forall i}{=} \sigma^2$$

i.e. $E[MSTR] = \sigma^2$ if $\tau_1 = \dots = \tau_k = 0$ is true.

5

For a general τ_i (may not be 0),

$$\bar{Y}_{i.} - \tau_i \stackrel{\text{ind.}}{\sim} N(\mu, \frac{\sigma^2}{b})$$

$$\bar{Y}_{..} = \frac{1}{k} \sum_{i=1}^k (\bar{Y}_{i.} - \tau_i) \sim N(\mu, \frac{\sigma^2}{kb})$$

Then $\frac{1}{k-1} E[SSTR]$

$$\begin{aligned} &= \frac{1}{k-1} E \left[b \cdot \sum_{i=1}^k (\bar{Y}_{i.} - \tau_i - \bar{Y}_{..} + \tau_i)^2 \right] \\ &= \frac{1}{k-1} E \left[b \cdot \sum_{i=1}^k (\bar{Y}_{i.} - \tau_i - \bar{Y}_{..})^2 \right] + \frac{1}{k-1} E \left[b \sum_{i=1}^k \tau_i^2 \right] \\ &\quad + E \left[\frac{2b}{k-1} \sum_{i=1}^k (\bar{Y}_{i.} - \tau_i - \bar{Y}_{..}) \cdot \tau_i \right] \end{aligned}$$

Note that τ_i is a parameter (constant) and

$$E(\bar{Y}_{i.} - \tau_i - \bar{Y}_{..}) = \mu - \mu = 0$$

$$\begin{aligned} \therefore E[MSTR] &= E \left[\frac{SSTR}{k-1} \right] \\ &= \sigma^2 + \frac{b}{k-1} \sum_{i=1}^k \tau_i^2 \end{aligned}$$

Similarly we can show that

$$E[MSB] = \sigma^2 + \frac{k}{b-1} \sum_{j=1}^b \beta_j^2 \quad (\text{Homework Problem})$$