

Section 7.3 General Factorial and 2^k Factorial Design

Cross Factorial Design with fixed factors

- Model (with three fixed factors, each with multiple levels)

$$Y_{ijkl} = \mu + \alpha_i + \beta_j + \gamma_l + (\alpha\beta\gamma)_{ij} + (\beta\gamma)_{jk} + (\alpha\gamma)_{il} + (\alpha\beta\gamma)_{ijk} + \varepsilon_{ijkl},$$

$$i = 1, \dots, a; j = 1, \dots, b; k = 1, \dots, c; l = 1, \dots, n$$

where iid errors $\varepsilon_{ijkl} \sim N(0, \sigma^2)$

- 2^k Factorial Design with each factor two levels

Table 7.3-3 Example of a 2^2 Factorial Experiment

Run	Design		x_1x_2	Average Yield	Individual Observations
	x_1	x_2			
1	-	-	+	55.0	55.5, 54.5
2	+	-	-	60.6	60.2, 61.0
3	-	+	-	64.2	64.5, 63.9
4	+	+	+	68.2	67.7, 68.7

$$\text{Average} = (55.0 + 60.6 + 64.2 + 68.2)/4 = 62.0$$

$$(1) = (-55.0 + 60.6 - 64.2 + 68.2)/4 = 2.4$$

$$(2) = (-55.0 - 60.6 + 64.2 + 68.2)/4 = 4.2$$

$$(12) = (55.0 - 60.6 - 64.2 + 68.2)/4 = -0.4$$

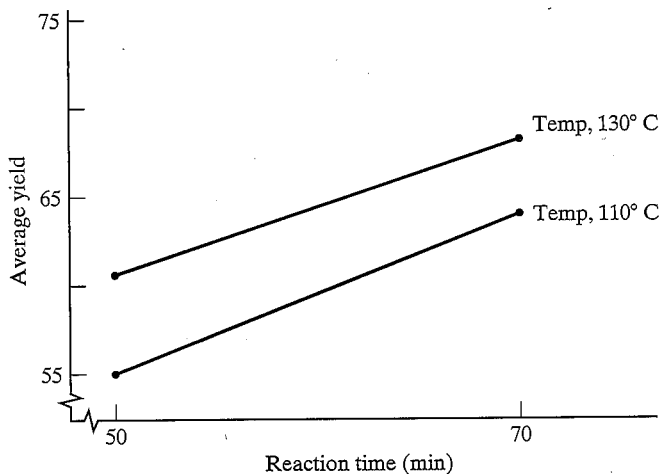


Table 7.3-4 The 2^3 Factorial Experiment

Run	Design							Observation
	x_1	x_2	x_3	x_1x_2	x_1x_3	x_2x_3	$x_1x_2x_3$	
1	-	-	-	+	+	+	-	Y_1
2	+	-	-	-	-	+	+	Y_2
3	-	+	-	-	+	-	+	Y_3
4	+	+	-	+	-	-	-	Y_4
5	-	-	+	+	-	-	+	Y_5
6	+	-	+	-	+	-	-	Y_6
7	-	+	+	-	-	+	-	Y_7
8	+	+	+	+	+	+	+	Y_8

The main effect of factor i ($i = 1, 2, 3$) is one-half the difference between the averages of the responses at the high and low levels of factor i . Thus, from the cube in Table 7.3-4, we find that

$$\begin{aligned}
 (1) &= \frac{1}{2} \left(\frac{Y_2 + Y_4 + Y_6 + Y_8}{4} - \frac{Y_1 + Y_3 + Y_5 + Y_7}{4} \right) \\
 &= \frac{-Y_1 + Y_2 - Y_3 + Y_4 - Y_5 + Y_6 - Y_7 + Y_8}{8},
 \end{aligned}$$

Example As an illustration of a 2^4 factorial, we use the data from an experiment designed to evaluate the effect of laundering on certain fire-retardant treatments for fabrics. **7.3-2** [M. G. Natrella, *Experimental Statistics*, National Bureau of Standards Handbook 91 (Washington, D.C.: U.S. Government Printing Office, 1963)]. Factor 1 is the type of fabric (sateen or monk's cloth), factor 2 corresponds to two different fire-retardant

Table 7.3-5 A 2^4 Factorial Experiment

x_1	x_2	x_3	x_4	Y	Effect
-	-	-	-	42	Average = $575/16 = 35.94$
+	-	-	-	31	(1) = $-129/16 = -8.06$
-	+	-	-	45	(2) = 1.56
+	+	-	-	29	(3) = -0.56
-	-	+	-	39	(4) = -0.56
+	-	+	-	28	(12) = -2.19
-	+	+	-	46	(13) = -0.31
+	+	+	-	32	(14) = -1.56
-	-	-	+	40	(23) = 0.81
+	-	-	+	30	(24) = 0.06
-	+	-	+	50	(34) = -0.31
+	+	-	+	25	(123) = 0.31
-	-	+	+	40	(124) = -1.19
+	-	+	+	25	(134) = -0.56
-	+	+	+	50	(234) = -0.44
+	+	+	+	23	(1234) = 0.06

For example, the $x_1x_2x_3$ column reads (written as a row)

- + + - + - - + - + + - + - - +.

Thus

$$\begin{aligned}
 & -42 + 31 + 45 - 29 + 39 - 28 - 46 + 32 \\
 (123) = & \frac{-40 + 30 + 50 - 25 + 40 - 25 - 50 + 23}{16} = 0.31.
 \end{aligned}$$

treatments, factor 3 describes the laundering condition (no laundering, after one laundering), and factor 4 corresponds to two different methods of conducting the flame test. The observations listed in Table 7.3-5 are in inches burned, measured on a standard-size sample fabric after a flame test.

TABLE 8.3-6 Yates Algorithm for the Data in Example 8.3-2

| x_1 | x_2 | x_3 | x_4 | y | <i>col 1</i> | <i>col 2</i> | <i>col 3</i> | <i>col 4</i> | Effect |
|-------|-------|-------|-------|-----|--------------|--------------|--------------|--------------|-----------------|
| — | — | — | — | 42 | 73 | 147 | 292 | 575 | Average = 35.94 |
| + | — | — | — | 31 | 74 | 145 | 283 | —129 | (1) = —8.06 |
| — | + | — | — | 45 | 67 | 145 | —52 | 25 | (2) = 1.56 |
| + | + | — | — | 29 | 78 | 138 | —77 | —35 | (12) = —2.19 |
| — | — | + | — | 39 | 70 | —27 | 12 | —9 | (3) = —0.56 |
| + | — | + | — | 28 | 75 | —25 | 13 | —5 | (13) = —0.31 |
| — | + | + | — | 46 | 65 | —35 | —8 | 13 | (23) = 0.81 |
| + | + | + | — | 32 | 73 | —42 | —27 | 5 | (123) = 0.31 |
| — | — | — | + | 40 | —11 | 1 | —2 | —9 | (4) = —0.56 |
| + | — | — | + | 30 | —16 | 11 | —7 | —25 | (14) = —1.56 |
| — | + | — | + | 50 | —11 | 5 | 2 | 1 | (24) = 0.06 |
| + | + | — | + | 25 | —14 | 8 | —7 | —19 | (124) = —1.19 |
| — | — | + | + | 40 | —10 | —5 | 10 | —5 | (34) = —0.31 |
| + | — | + | + | 25 | —25 | —3 | 3 | —9 | (134) = —0.56 |
| — | + | + | + | 50 | —15 | —15 | 2 | —7 | (234) = —0.44 |
| + | + | + | + | 23 | —27 | —12 | 3 | 1 | (1234) = 0.06 |

Let us assume that there are n independent observations $Y_{i1}, Y_{i2}, \dots, Y_{in}$ with variance estimate $S_i^2 = [\sum_{j=1}^n (Y_{ij} - \bar{Y}_i)^2]/(n-1)$ at each of the 2^k level combinations, $i = 1, 2, \dots, 2^k$. The 2^k variance estimates can be pooled to obtain the overall variance estimate

$$S^2 = \frac{1}{2^k} \sum_{i=1}^{2^k} S_i^2 = \frac{1}{(n-1)2^k} \sum_{i=1}^{2^k} \sum_{j=1}^n (Y_{ij} - \bar{Y}_i)^2.$$

Because the estimate of the variance of an average $\bar{Y}_i$ at a particular level combination is S^2/n , and because the overall average and each estimated effect can be written as $(1/2^k) \sum_{i=1}^{2^k} c_i \bar{Y}_i$, where the coefficients c_i are either $+1$ or -1 , we find that the estimate of the variance of an effect is

$$\text{var}(\text{effect}) = \text{var}(\text{average}) = \frac{1}{(2^k)^2} \sum_{i=1}^{2^k} \text{var}(\bar{Y}_i) = \frac{S^2}{n2^k}.$$

The estimated effects, together with estimates of their standard deviations which are also known as standard errors, indicate the statistical significance of the various effects.

Example
7.3-3

Using the result derived in Exercise 7.3-1, we find for the data in Table 7.3-3 that

$$\begin{aligned}s^2 &= \frac{1}{4} \left[\frac{(55.5 - 54.5)^2}{2} + \frac{(60.2 - 61.0)^2}{2} + \frac{(64.5 - 63.9)^2}{2} + \frac{(67.7 - 68.7)^2}{2} \right] \\ &= 0.375,\end{aligned}$$

and

$$\text{var}(\text{effect}) = \text{var}(\text{average}) = \frac{1}{(2)(4)}(0.375) = 0.0469.$$

Hence, the standard error of an estimated effect, as well as of the average, is

$$[\text{var}(\text{effect})]^{1/2} = [\text{var}(\text{average})]^{1/2} = 0.22.$$

Thus, the two-sigma limits around the estimates are 62.0 ± 0.44 for the mean, 2.4 ± 0.44 for the main effect of factor 1, 4.2 ± 0.44 for the main effect of factor 2, and -0.40 ± 0.44 for the two-factor interaction. These intervals are approximate 95 percent confidence intervals and indicate large main effects, but negligible interaction.

Table 7.3-6 Normal Scores of the $m = 2^{4-1} = 15$ Estimated Effects from Example 7.3-2

| Identity of Effect | Effect by Magnitude | Rank | $P_i = \frac{i - 0.5}{m}$ | z_i |
|--------------------|---------------------|------|---------------------------|-------|
| (1) | -8.06 | 1 | 0.033 | -1.84 |
| (12) | -2.19 | 2 | 0.100 | -1.28 |
| (14) | -1.56 | 3 | 0.167 | -0.97 |
| (124) | -1.19 | 4 | 0.233 | -0.73 |
| (3) | -0.56 | 6 | 0.367 | -0.34 |
| (4) | -0.56 | 6 | 0.367 | -0.34 |
| (134) | -0.56 | 6 | 0.367 | -0.34 |
| (234) | -0.44 | 8 | 0.500 | 0.00 |
| (13) | -0.31 | 9.5 | 0.600 | 0.25 |
| (34) | -0.31 | 9.5 | 0.600 | 0.25 |
| (24) | 0.06 | 11.5 | 0.733 | 0.62 |
| (1234) | 0.06 | 11.5 | 0.733 | 0.62 |
| (123) | 0.31 | 13 | 0.833 | 0.97 |
| (23) | 0.81 | 14 | 0.900 | 1.28 |
| (2) | 1.56 | 15 | 0.967 | 1.84 |

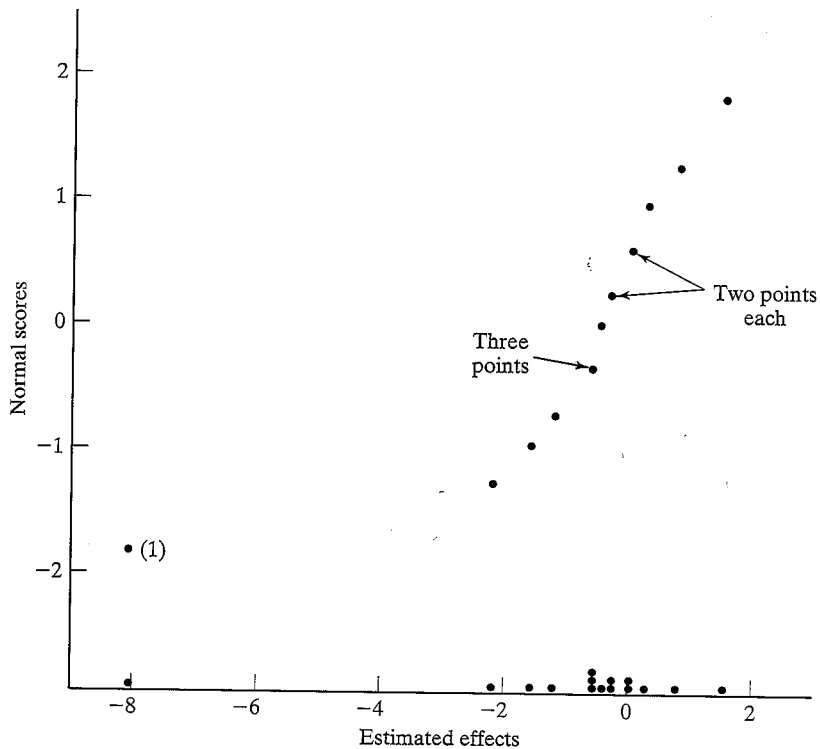


Figure 7.3-1 Dot diagram and normal probability plot of the estimated effects from Example 7.3-2

TABLE 8.4-1 The 2^{3-1} Fractional Factorial with $x_3 = x_1x_2$

| Run | Design | | | Y |
|-----|--------|-------|-------|-------|
| | x_1 | x_2 | x_3 | |
| 1 | — | — | + | Y_1 |
| 2 | + | — | — | Y_2 |
| 3 | — | + | — | Y_3 |
| 4 | + | + | + | Y_4 |

$$L_0 = (Y_1 + Y_2 + Y_3 + Y_4)/4 \rightarrow \mu + (123)$$

$$L_1 = (-Y_1 + Y_2 - Y_3 + Y_4)/4 \rightarrow (1) + (23)$$

$$L_2 = (-Y_1 - Y_2 + Y_3 + Y_4)/4 \rightarrow (2) + (13)$$

$$L_3 = (Y_1 - Y_2 - Y_3 + Y_4)/4 \rightarrow (3) + (12)$$