Section 6.3 Randomized Complete Block Design (RCBD)

Examples of RCBD

- A chemist is studying the reaction rate of five chemical agents. Only five agents can be analyzed effectively per day. Since day-to-day differences may affect the reaction rate, each day is used as a block. All five chemical agents are tested each day in independently random orders.
- <u>Blocking</u> is to arrange experimental units in groups that are similar to each other.
- Blocking factor normally is not of primary interest to the investigator, mainly to remove nuisance effects.
- <u>Blocking factor</u> is an observational factor, related to the characteristic of the experimental unit, or the experimental settings. Not an experimental factor.

TABLE 7.3-1 Compressive Strength of Concrete (100 pounds per square inch)

| | Batch | | | | | |
|------------|-------|----|----|----|----|----------------|
| Treatment | 1 | 2 | 3 | 4 | 5 | Treatment Mean |
| A | 52 | 47 | 44 | 51 | 42 | 47.2 |
| В | 60 | 55 | 49 | 52 | 43 | 51.8 |
| C | 0.000 | 48 | | | | 46.2 |
| Batch mean | 56 | 50 | 46 | 49 | 41 | 48.4 |

TABLE 7.3-2 Data from a Randomized Complete Block Experiment

| | Block | | | | | | |
|----------------------------|--|--|-----|------------------------------------|--|--|--|
| Treatment | 1 | 2 | | j | | b | Treatment Mean |
| 1 2 : : : k | $Y_{11} \\ Y_{21} \\ \vdots \\ X_{k1}$ | $Y_{12} \\ Y_{22} \\ \vdots \\ Y_{k2}$ | ••• | : : Y _{ij} : | | Y_{1b} Y_{2b} \vdots \vdots Y_{kb} | $egin{array}{c} ar{Y}_1 \ ar{Y}_2 \ ar{Y}_i \ ar{Y}_i \ ar{Y}_k \ \end{array}$ |
| Block mean | Ῡ., | <u><u> </u></u> | | $ar{Y}_{\cdot j}$ | | <u> </u> Т. _ь | Ÿ |

RCBD with fixed factor and fixed block factor

Model (with k treatments and b blocks)

$$Y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}, i = 1, ..., k; j = 1, ..., b.$$
 where iid $\varepsilon_{ij} \sim N(0, \sigma^2), \sum_{i=1}^k \tau_i = \sum_{j=1}^b \beta_j = 0.$

SS Decomposition: SSTO = SSTR + SSB+SSE

$$SSTO = \sum_{i=1}^{k} \sum_{j=1}^{b} (Y_{ij} - \overline{Y}_{..})^{2}, \overline{Y}_{..} = \frac{1}{kb} \sum_{i=1}^{k} \sum_{j=1}^{b} Y_{ij};$$

$$SSTR = \sum_{i=1}^{k} b (\overline{Y}_{i.} - \overline{Y}_{..})^{2}, \overline{Y}_{i.} = \frac{1}{b} \sum_{j=1}^{b} Y_{ij}, i = 1,..., k;$$

$$SSB = \sum_{j=1}^{b} k (\overline{Y}_{.j} - \overline{Y}_{..})^{2}, \overline{Y}_{.j} = \frac{1}{k} \sum_{i=1}^{k} Y_{ij}, j = 1,..., b;$$

$$SSE = \sum_{i=1}^{k} \sum_{j=1}^{b} (Y_{ij} - \overline{Y}_{i.} - \overline{Y}_{.j} + \overline{Y}_{..})^{2}$$

ANOVA for RCBD

| Source of Variation | SS | DF | MS | F |
|---------------------|------|------------|------------------------|----------------------------|
| Treatment | SSTR | k-1 | MSTR = SSTR/(k-1) | FTR = MSTR / MSE |
| Block | SSB | b-1 | MSB = SSB/(b-1) | F _B = MSB / MSE |
| Error | SSE | (k-1)(b-1) | MSE = SSE/((k-1)(b-1)) | |
| Total | SSTO | kb-1 | | |

Distribution of F-statistic in ANOVA

$$E(MSE) = \sigma^2, E(MSTR) = \sigma^2 + \frac{b}{k-1} \sum_{i=1}^k \tau_i^2,$$

$$E(MSB) = \sigma^2 + \frac{k}{b-1} \sum_{j=1}^{b} \beta_j^2.$$

By Cochran's Theorem, SSTR, SSB, and SSE are mutually independent.

Under $H_0: \tau_i = 0, i = 1,..., k, (H_1: at least one \tau_i \neq 0)$

$$F_{TR} = \frac{SSTR/(k-1)}{SSE/((k-1)(b-1))} \sim F(k-1,(k-1)(b-1)).$$

Under $H_0: \beta_j = 0, j = 1,..., b, (H_1: at least one \beta_j \neq 0)$

$$F_B = \frac{SSB/(b-1)}{SSE/((k-1)(b-1))} \sim F(b-1,(k-1)(b-1)).$$

Parameter Estimation and Comparison Test

ANOVA model

LSE or MLE: $\hat{\mu} = \overline{Y}$,

$$\hat{\tau}_i = \overline{Y}_{i.} - \overline{Y}_{..}, \hat{\beta}_j = \overline{Y}_{.j} - \overline{Y}_{..}, i = 1,...k, j = 1,....b.$$

Reference Distribution for \overline{Y}_i , i = 1,...,k:

$$\overline{Y}_{i.} \sim N\left(\mu + \tau_{i}, \frac{\sigma^{2}}{b}\right) \Rightarrow \frac{\overline{Y}_{i.} - (\mu + \tau_{i})}{\sqrt{\frac{MSE}{b}}} \sim t((k-1)(b-1)), i = 1, ..., k$$

= 100(1- α)% Confidence Interval of i-th treatment mean($\mu + \tau_i$):

$$\overline{Y}_{i.} \pm t_{\frac{\alpha}{2}}((k-1)(b-1)) \cdot \sqrt{\frac{MSE}{b}}$$

Follow-up Comparison and Test

■ 100(1-α)% Confidence Interval of j-th block mean $(μ + β_i)$

$$\overline{Y}_{.j} \pm t_{\frac{\alpha}{2}}((k-1)(b-1)) \cdot \sqrt{\frac{MSE}{k}}$$

100(1-α)% Confidence Interval of (μ_i – μ_i) :

$$(\overline{Y}_{i.} - \overline{Y}_{j.}) \pm t_{\frac{\alpha}{2}}((k-1)(b-1)) \cdot \sqrt{\frac{2MSE}{b}}$$

100(1-α)% Tukey's Simultaneous C.I. for (μ_i – μ_j) any pair (i,j)

$$(\overline{Y}_{i.} - \overline{Y}_{j.}) \pm q(\alpha, k, (k-1)(b-1)) \cdot \sqrt{\frac{MSE}{b}}$$

Model Diagnosis

Normality Check for iid errors -- histogram and QQ plot

$$\varepsilon_{ij} \sim N(0, \sigma^2), \sum_{i=1}^k \tau_i = \sum_{j=1}^b \beta_j = 0.$$

 Variances in different treatment groups and block groups are roughly the same

$$\overline{Y}_{i.} = \frac{1}{b} \sum_{j=1}^{b} Y_{ij} \sim N \left(\mu + \tau_i, \frac{\sigma^2}{b} \right),$$

$$\overline{Y}_{.j} = \frac{1}{k} \sum_{i=1}^{k} Y_{ij} \sim N \left(\mu + \beta_{j}, \frac{\sigma^{2}}{k} \right)$$

Residual vs fitted value $\sum_{i=1}^k \sum_{j=1}^b e_{ij} \hat{Y}_{ij} = 0$

Example: (SAS code)

avgain is the mean average daily gain of all calves in a pen. Average daily gain is (final weight minus initial weight)/(number of days) on study.

| data weightgain; input farm pen diet \$ avgain; cards; 11C2.19 12D2.44 13B3.02 621A1.73 622D1.49 623B1.96 624C1.33; | /* Without Blcok Effect*/ proc glm data=weightgain; class diet; model avgain=diet; lsmeans diet / pdiff; run; /* with block effect */ proc glm data=weightgain; class farm diet; model avgain=farm diet / p clm; lsmeans diet / pdiff cl adjust=tukey; output out=OutData p=predict r=residual; run; |
|--|---|
| | proc gplot data=Outdata; plot predict*residual avgain*residual; run; |
| | |

Section 6.4 Design with two blocks

 Latin Square: each of k treatments is assigned to each level of each of the two blocking factors in a special way

| 1 | 2 | 3 | 4 |
|---|---|---|---|
| 2 | 1 | 4 | 3 |
| 3 | 4 | 1 | 2 |
| 4 | Э | 2 | 1 |

| $\lceil 1 \rceil$ | 2 | 3 | 4 | 5 | |
|-------------------|---|---|---|---|--|
| 2 | 4 | 1 | 5 | 3 | |
| 3 | 5 | 4 | 2 | 1 | |
| 4 | 1 | 5 | 3 | 2 | |
| 5 | 3 | 2 | 1 | 4 | |

ANOVA for Latin Square with two blocks

| Source of Variation | SS | DF | MS | F |
|---------------------|-------|----------|-------------------|--|
| Treatment | SSTR | k-1 | MSTR | F _{TR} = MSTR / MSE |
| Row | SSRow | k-1 | MS _{Row} | F _{Row} = MS _{Row} / MSE |
| Column | SScol | k-1 | MScol | F _{Col} = MS _{Col} / MSE |
| Error | SSE | k^2-3k+2 | MSE | |
| Total | SSTO | k^2-1 | | |

Example 4-1

The yields (in bushels per 0.1 of an acre) of 4 corn hybrids.

| | | Column | | | Row Total | Treatment Total |
|-----------------|------|--------|------|------|-----------|---------------------|
| | 10-A | 14-B | 7-C | 8-D | 39 | $T_A = 53$ |
| Row | 7-D | 18-A | 11-B | 8-C | 44 | T _B = 44 |
| | 5-C | 10-D | 11-A | 9-B | 35 | $T_{\rm C} = 30$ |
| | 10-B | 10-C | 12-D | 14-A | 46 | $T_D = 37$ |
| Column Total | 32 | 52 | 41 | 39 | | |

$$H_0$$
: $T_A = T_B = T_C = T_D = 0$

$$k-1=3$$
, $k^2-3k+2=16-12+2=6$, $F_{0.01}(3, 6)=9.78$, p-value<0.01.